Adjunctive Risperidone for Schizophrenia: Meta-analysis of Randomized Double-blind, Placebo-controlled Trials

Authors

Wei Zheng^{1*}, Xian-Bin Li^{2*}, Zhan-Ming Shi^{3*}, Xin-Hu Yang¹, Dong-Bin Cai⁴, Chee H. Ng⁵, Gabor S. Ungvari^{6,7}, Wei-Jian Liu¹, Yu-Jie Wu¹, Yuan-Yuan Wang⁸, Yu-Ping Ning¹, Yu-Tao Xiang⁹

Affiliations

- 1 The Affiliated Brain Hospital of Guangzhou Medical University (Guangzhou Huiai Hospital), Guangzhou, China
- 2 The National Clinical Research Center for Mental Disorders & Beijing Key Laboratory of Mental Disorders, Beijing Anding Hospital & the Advanced Innovation Center for Human Brain Protection, Capital Medical University, Beijing, China
- 3 Chongqing Jiangbei Mental Health Center, Chongqing, China
- 4 Shenzhen Traditional Chinese Medicine Hospital, Shenzhen, China
- 5 Department of Psychiatry, University of Melbourne, Melbourne, Victoria, Australia
- 6 The University of Notre Dame Australia, Fremantle, Australia
- 7 Division of Psychiatry, School of Medicine, University of Western Australia, Perth, Australia
- 8 Faculty of Health and Life Sciences, De Montfort University, Leicester, United Kingdom
- 9 Unit of Psychiatry, Institute of Translational Medicine, Faculty of Health Sciences, University of Macau, Macao SAR, China

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Correspondence

Dr. Yu-Tao Xiang
Faculty of Health Sciences,
University of Macau,
3/F, Building E12,
Avenida da Universidade,
Taipa,
Macao SAR,
China
xyutly@gmail.com

Dr. Yu-Ping Ning
The Affiliated Brain Hospital of Guangzhou Medical
University (Guangzhou Huiai Hospital),
Guangzhou,
China
ningjeny@126.com

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ABSTRACT

Background Results of previous studies on the safety and efficacy of adjunctive risperidone for schizophrenia have been inconsistent.

Aim The aim of this study was to examine the efficacy and tolerability of risperidone as an adjunct medication to antipsychotic treatment in a meta-analysis of randomized controlled trials (RCTs).

Methods Two independent investigators extracted data for a random effects meta-analysis and assessed the quality of studies using risk of bias and the Jadad scale. Weighted and standardized mean differences (WMDs/SMDs) and risk ratio (RR) ± 95 % confidence intervals (CIs) were calculated.

Results Nine RCTs (n = 630) with double-blind design were identified. Risperidone outperformed placebo in improving negative (9 RCTs, n = 602, SMD: −0.47 [95 % CI: −0.87, −0.07], p = 0.02; I² = 82 %), but not the overall, positive, and general psychopathology scores. The significant therapeutic effect on negative symptoms disappeared in the sensitivity analysis after removing an outlying study and in 50% (6/12) of the subgroup analyses. Risperidone outperformed placebo in reducing weight (3 RCTs, n = 186, WMD: −3.83 kg, p = 0.04; I² = 92 %) and body mass index (WMD: −2.23 kg/m², p = 0.04; I² = 95 %). Risperidone

* These authors contributed equally to the paper.

caused dry mouth but was associated with less weight gain overall and weight gain of $\geq 7\%$ of the initial weight. All-cause discontinuation and other adverse events were similar between reboxetine and placebo.

Conclusion Adjunctive reboxetine could be useful for attenuating antipsychotic-induced weight gain, but it was not effective in treating psychopathology including negative symptoms in schizophrenia.

Introduction

Schizophrenia is a severe and chronic psychiatric disorder. Nearly 50 % of schizophrenia patients fail to respond to antipsychotic (AP) monotherapy, particularly in terms of negative symptoms [1–3]. Augmentation strategies are commonly used to enhance AP efficacy and to reduce adverse drug reactions (ADRs) induced by APs [2, 4–6].

Most negative symptoms (e. g., anhedonia and social withdrawal) are likely to be intrinsic to schizophrenia [7] and associated with neurocognitive deficits [8] including impaired attention, memory, and executive functions [9]. Certain augmentation strategies appear to be effective to improve the negative symptoms of schizophrenia; for example, adjunctive antidepressants have a moderate effect size for negative symptoms [7, 10].

In the past two decades, weight gain induced by APs has attracted increased attention [11–16]. Weight gain is associated with poor treatment adherence and quality of life, higher rate of medical comorbidity particularly cardiovascular diseases, diabetes mellitus, and osteoarthritis and mortality [12, 13, 17–21]. Reboxetine, a norepinephrine reuptake inhibitor, is an antidepressant and an anxiety drug [18, 22]. According to a recent meta-analysis, as an antidepressant, reboxetine is ineffective and causes side effects [23]; however, in another network meta-analysis, reboxetine appeared to have antidepressant effect in major depression [24]. Some trials have found that reboxetine is safe and efficacious in treating negative symptoms of schizophrenia [25–28] and AP-induced weight gain in some [18, 22, 27] but not all [29, 30] randomized controlled trials (RCTs).

The efficacy of reboxetine in treating negative symptoms and/or reducing AP-induced weight gain in schizophrenia have been examined in meta-analyses [10, 12, 31–34]. Common limitations of these meta-analyses include the small number of included studies, resulting in insufficient power. For example, 3 meta-analyses [12, 33, 34], each with 2 RCTs ($n = 85$) [18, 22], found that reboxetine was superior to placebo in reducing AP-induced weight gain. Another meta-analysis [31] included 9 RCTs, but one of them was an open-label study [35], the inclusion of which violated standard recommendations [36]. In addition, non-English databases were not searched for relevant RCTs [26].

The aim of this study was to obtain more robust evidence regarding the efficacy and safety of adjunctive reboxetine added to APs. To this end, a comprehensive meta-analysis was conducted involving all RCTs on reboxetine added to all types of APs in treating schizophrenia including also recent RCTs published in Chinese, which were not included in previous meta-analyses.

Methods

Types of studies

According to the PICOS acronym, the selection criteria were as follows: participants (P), patients with schizophrenia diagnosed according to any criteria; intervention (I), reboxetine plus APs; comparison (C), APs plus placebo; outcomes (O), efficacy and safety of adjunctive reboxetine with meta-analyzable data; study design (S), randomized, double-blind, placebo-controlled trials. A methodologically sound RCT [37] was excluded because patients were selected from the combined sample of 2 other RCTs [18, 22].

Outcome measures

Clinical outcomes were recorded based on intent-to-treat (ITT) analysis, if provided. The co-primary outcome measures were the change of negative symptoms assessed with the Positive and Negative Syndrome Scale (PANSS) [38] or the Brief Psychiatric Rating Scale (BPRS) [39] or the total scores of the Scale for the Assessment of Negative Symptoms (SANS) [40] and body weight (kg). Key secondary outcomes were the changes of total, positive, and general psychopathology scores of the PANSS or BPRS or the total scores of the Scale for the Assessment of Positive Symptoms (SAPS) [41], body mass index (BMI, kg/m^2), cognitive functions, ADRs, and all-cause discontinuation rate.

Study selection

The PubMed, PsycINFO, EMBASE, the Cochrane Library, Chinese Journal Net, WanFang, and the China Biology Medicine databases were independently and systematically searched by 2 reviewers from their inception until November 3, 2016. The keywords for the search were (reboxetine OR Edronax) AND (schizophrenic disorder OR disorder, schizophrenic OR schizophrenic disorders OR schizophrenia OR dementia praecox). Reference lists from review articles [10, 12, 31–34] were hand-searched for additional studies. First/ corresponding authors were contacted for missing information, whenever necessary.

Data extraction

Data were independently identified, checked, extracted, and analyzed by 2 reviewers. Inconsistencies were resolved by consensus involving a third reviewer. If data from the same study were reported in more than 1 RCT, only the RCT with complete data was included in the analyses.

Statistical methods

According to the guidelines of the Preferred Reporting Items for Systematic Reviews and Meta-Analyses (PRISMA) statement [42], data

were summarized statistically using the Review Manager version 5.3 (<http://www.cochrane.org>) if they were available and sufficiently similar. A random effects model was used in all cases [43]. For the meta-analytic pooling of continuous outcomes, weighted or standardized mean differences (WMDs or SMDs) with their 95 % confidence intervals (CIs) are reported. Summary statistics of dichotomous outcomes are presented as risk ratios (RRs) $\pm$ 95 % CIs. When RRs were significant, number-needed-to-harm (NNH) was calculated by dividing 1 by the risk difference. Whenever both change score and endpoint values of a continuous outcome were available, change scores were preferred. Missing standard deviation (SD) was replaced by the average SD of other RCTs following the suggestion of Leucht et al. [36]. In cases of $I^2 > 50\%$ for co-primary outcomes, reasons were sought to explain the heterogeneity by conducting a sensitivity analysis (i. e., removing 1 outlier [SMD ≤ -1.3] study) [27]. In addition, the following 6 subgroup analyses were performed to identify the reasons for the heterogeneity of significance: (1) Chinese versus non-Chinese studies; (2) clozapine vs. other APs; (3) trial duration (weeks): ≥ 12 vs. < 12 (mean splitting method for trial duration); (4) age: ≥ 38.0 vs. < 38.0 years (mean splitting method for age); (5) male predominance ($\geq 60\%$) vs. no sex predominance; (6) study quality: Jadad score ≥ 3 vs. < 3 . The above subgroup analyses were repeated after leaving out 1 outlying study [27]. Funnel plots and Egger's test [44] were used to

judge publication bias. All analyses were 2 tailed, with alpha set at 0.05.

Assessment of the studies

Cochrane risk of bias [45] was employed to assess the methodological quality of RCTs (► **Fig. 15**). The Grading of Recommendations, Assessment, Development, and Evaluation (GRADE) system [46, 47] was performed to estimate the recommendation for outcome measures of adjunctive reboxetine for schizophrenia. Following the methodology of other studies [31, 48], the Jadad scale (range: 0–5) assessed the quality of included studies (► **Table 1**) [49]. The Jadad total score of ≥ 3 indicated “high quality” [50].

Results

Results of the search

The original search from the above databases yielded 339 electronic records and 1 trial retrieved by hand-search (► **Fig. 1**). By the end of screening all papers, 9 RCTs [18, 22, 25–30, 51] published in English (6 RCTs) [18, 22, 28–30, 51] and Chinese (3 RCTs, (► **Table 15**) [25–27] were eligible and analyzed.

► **Fig. 1** PRISMA flow diagram.

Study characteristics

All 9 RCTs ($n = 630$) were double-blind and 4 (44.4%, $n = 211$) used ITT analyses. The mean study duration was 11.8 ± 7.6 (range: 4–24, median: 12.0) weeks (► **Table 1**). Three RCTs were conducted in China ($n = 363$), 2 in Israel ($n = 85$), and 1 each in Iran ($n = 50$), South Africa ($n = 30$), Germany ($n = 35$), and Spain ($n = 67$).

Patient characteristics

The mean age of the 630 patients was 38.0 ± 6.9 years (range: 29.9–50.4, median: 40.4 years); males accounted for $79.5 \pm 18.9\%$ (range: 50.8–100, median: 73.1) of the sample. The mean illness duration was 10.6 ± 9.4 (range: 0.4–27.2, median: 8.5) years (8 RCTs with available data) (► **Table 1**). Six RCTs were conducted in inpatients ($n = 393$), 1 RCT involved both in- and outpatients ($n = 135$), and the type of patients in 2 RCTs ($n = 102$) was unspecified.

Treatment characteristics

Reboxetine dose was 6.2 ± 1.8 mg/day (range: 4.0–8.0, median: 6.0 mg/day). Baseline APs included olanzapine (2 RCTs, $n = 85$), clozapine (2 RCTs, $n = 228$), haloperidol (3 RCTs, $n = 215$), and multiple APs (2 RCT, $n = 102$) (► **Table 1**).

Quality assessment

While 8 RCTs (88.9%) [18, 22, 25–29, 51] were rated as low risk regarding attrition bias, 4 RCTs [18, 22, 29, 30] described an adequate method of random sequence generation and 2 RCTs [29, 30] were rated as low risk regarding the allocation concealment methods (► **Fig. 1S**). In addition, 5 RCTs (63%) [18, 22, 28, 29, 51] were rated as low risk for selective report bias. According to the GRADE approach, the quality of evidence for 10 outcomes ranged from “low” (20%) via “moderate” (40%) to “high” (40%) (► **Table 2S**). The mean score of the Jadad scale was 3.6 ± 1.1 (range: 2–5; median: 3.0); 8 RCTs (88.9%) were classified as high quality (► **Table 1**).

Psychotic symptoms

Meta-analysis of negative symptoms (PANSS [6 RCTs] and SANS [3 RCTs]) showed that adjunctive reboxetine was superior to placebo (9 RCTs, $n = 602$, SMD: -0.47 [95% CI: $-0.87, -0.07$], $p = 0.02$; $I^2 = 82\%$, ► **Fig. 2**). The significance (SMD: -0.36 [95% CI: $-0.77, 0.05$], $p = 0.09$; $I^2 = 79\%$) disappeared after 1 outlying (SMD ≤ -1.3) study [27] was removed. In subgroup analyses, the significance also disappeared in 6 out of the 12 subgroups (► **Table 2**). The superiority of reboxetine disappeared in non-Chinese studies ($p = 0.59$), with APs other than clozapine ($p = 0.27$), trials lasting less than 12 weeks ($p = 0.37$), male predominance ($\geq 60\%$) ($p = 0.10$), mean age younger than 38.0 years ($p = 0.54$), and having a Jadad score more than 3 ($p = 0.10$). (**Table 2S**) presents subgroup analyses after leaving out 1 outlying study [27] and found that significance disappeared in 7 out of the 12 subgroups. The superiority of reboxetine disappeared in non-Chinese studies ($p = 0.59$), APs other than clozapine ($p = 0.27$), trials lasting less than 12 weeks ($p = 0.37$), male predominance ($\geq 60\%$) ($p = 0.26$), mean age younger ($p = 0.54$) or older than 38.0 years ($p = 0.11$), and having a Jadad score more than 3 ($p = 0.26$).

No group difference was found in change of overall psychopathology measured with the PANSS (6 RCTs) ($n = 473$, SMD: -0.50

[95% CI: $-1.05, 0.06$], $p = 0.08$; $I^2 = 88\%$; ► **Fig. 2**), positive symptoms (PANSS [6 RCTs] and SAPS [3 RCTs]) ($n = 602$, SMD: -0.00 [95% CI: $-0.16, 0.16$], $p = 0.98$; $I^2 = 0\%$; ► **Fig. 2**) and PANSS general psychopathology score (5 RCTs) ($n = 438$, SMD: -0.46 [95% CI: $-0.97, 0.05$], $p = 0.08$; $I^2 = 85\%$, ► **Fig. 2**).

Weight change

Compared to placebo, reboxetine caused significant weight (3 RCTs, $n = 186$, WMD: -3.83 kg [95% CI: $-7.40, -0.26$], $p = 0.04$; $I^2 = 92\%$; ► **Fig. 3**) and BMI reduction (3 RCTs, $n = 186$, WMD: -2.23 kg/m² [95% CI: $-4.35, -0.12$], $p = 0.04$; $I^2 = 95\%$, ► **Fig. 3**). The results concerning weight (WMD: -1.90 kg [95% CI: $-3.07, -0.72$], $p = 0.002$; $I^2 = 0\%$) was consistent even after 1 outlier (SMD < -1.0) study [27] was removed.

Cognitive functions

Only 2 RCTs [26, 27] assessed cognitive functions. One study [26] found reboxetine superior to placebo in memory quotient, recognition, and associative learning assessed by the Wechsler Memory Scale-Revised, Chinese version. The other study [27] found reboxetine outperformed placebo in attention, immediate memory, and delayed memory assessed with the Repeatable Battery for the Assessment of Neuropsychological Status. Because of the different scales used, meta-analysis of cognitive functions was not possible.

Discontinuation rate and ADRs

All-cause discontinuations were similar between reboxetine and placebo (8 RCTs, $n = 580$, RR: 1.05 [95% CI: 0.71, 1.56], $p = 0.81$, $I^2 = 0\%$; ► **Fig. 2S**). Regarding ADRs, reboxetine caused more frequent dry mouth ($p = 0.04$, NNH = 14, 95% CI: 7–50; ► **Fig. 3S**) but was associated with less weight gain overall ($p = 0.01$, NNH = 7, 95% CI: 4–50) and weight gain of $\geq 7\%$ of the initial weight ($p = 0.006$, NNH = 3, 95% CI: 2–8; ► **Fig. 3S**). Meta-analyses of akathisia, dizziness, insomnia, tachycardia, constipation, and nausea/vomiting showed no significant differences between reboxetine and placebo ($p = 0.21$ – 0.82 ; ► **Fig. 3S**).

Publication bias

Since a minimum of 10 RCTs are needed to conduct funnel plot or Egger's test [52], publication bias was not assessed for negative symptoms (9 RCTs) and weight change (3 RCTs).

Discussion

This meta-analysis found that reboxetine was not consistently effective in treating negative symptoms because the significant improvement was driven by an outlying study [27]. This result is consistent with previous findings [10, 32]. In this meta-analysis, adjunctive reboxetine was effective in weight reduction in the treatment of schizophrenia. Compared with Helfer et al.'s meta-analysis [32] of 5 RCTs [18, 22, 29, 30, 51], 4 additional RCTs [25–28] were included in this study, generating larger power and allowing more comprehensive analyses. A recent meta-analysis [31] concluded that reboxetine may improve negative symptoms, but the evidence presented was of a very low quality, probably due to the inclusion of an open-label study [35]. In the present meta-analysis, a significant effect of reboxetine on negative symptoms was only

► **Table 1** Characteristics of the studies included in the meta-analysis.

Study (country)	Number of patients	Blinding	Analyses	Trial duration (weeks)	Setting (%)	Diagnosis (%)	Diagnostic criteria	Illness severity (PANSS/BPRS) ^a /duration ^a	Age ^a : years (range)	Sex ^a : male (%)	Placebo group: dose (mg/d): mean (range)	Reboxetine group: dose (mg/d): mean (range)	Risk of Bias [*]	Jadad score
Hinkelmann et al. 2013 (Germany)	T: 35 M; 16 A; 19	DB	ITT	4	NR	Sz (100)	DSM-IV	-32.9; -12.6 years	40.4 (NR)	68.6	APs: $\emptyset$ = NR (NR)	REB: $\emptyset$ = NR (4-8)	2	3
Li et al. 2008 (China)	T: 118 M; 59 A; 59	DB	OC	12	Inpatients (100)	Sz (100)	CCMD-3	-75.2; -27.2 years	50.4 (18-65)	100	CLZ: $\emptyset$ = 282.5 (NR)	REB: $\emptyset$ = 8 ^b (4-8)	2	3
Poyurovsky et al. 2003 (Israel)	T: 26 M; 13 A; 13	DB	OC	6	Inpatients (100)	Sz (100)	DSM-IV	-NR; -3.0 years	30.8 (19-49)	65.4	OLA: $\emptyset$ = 10 (FD)	REB: $\emptyset$ = 4 (FD)	5	5
Poyurovsky et al. 2007 (Israel)	T: 59 M; 28 A; 31	DB	ITT	6	Inpatients (100)	Sz (NR), SzD (NR)	DSM-IV	-NR; -3.5 years	29.9 (19-48)	64.4	OLA: $\emptyset$ = 10 (FD)	REB: $\emptyset$ = 4 (FD)	5	5
Schutz et al. 2001 (South Africa)	T: 30 M; 15 A; 15	DB	NR	6	Inpatients (100)	Sz (100)	DSM-IV	-78.9; -8.1 years	32.5 (NR)	93.3	HAL: $\emptyset$ = 5 (FD)	REB: $\emptyset$ = 8 (FD)	4	3
Shafiti et al. 2015 (Iran)	T: 50 M; 25 A; 25	DB	ITT	12	Inpatients (100)	Sz (100)	DSM-IV-TR	-NR; -8.9 years	40.4 (NR)	100	HAL: $\emptyset$ = NR (5-20)	REB: $\emptyset$ = 4 (FD)	4	3
Sun et al. 2011 (China)	T: 135 M; 66 A; 69	DB	OC	12	In- and outpatients (NR)	Sz (100)	DSM-IV	-85.6; -0.4 years	32.5 (18-45)	50.8	HAL: $\emptyset$ = 14 (10-30)	REB: $\emptyset$ = 6 (4-12)	2	2
Usall et al. 2014 (Spain)	T: 67 M; 33 A; 34	DB	ITT	24	NR	Sz (100)	DSM-IV	-73.5; -NR	42.1 (18-65)	73.1	OLA/RIS: $\emptyset$ = NR (0.5-20) ^c	REB: $\emptyset$ = 8 ^b (4-8)	5	5
Zhao et al. 2013 (China)	T: 110 M; 55 A; 55	DB	OC	24	Inpatients (100)	Sz (100)	ICD-10	-73.6; -21.4 years	43.1 (18-55)	100	CLZ: $\emptyset$ = 302.3 (NR)	REB: $\emptyset$ = 7.6 (4-8)	3	3

* Number of low risk judgements; $\emptyset$ = mean; ^a Data were extracted from mean baseline values of each studies; ^b The active drug of these trials started with a low dose and reached the target dosage within 1 week; ^c Doses of antipsychotic drugs are expressed in risperidone equivalent doses; A: augmentation; BPRS: Brief Psychiatric Rating Scale; CCMD-3: China's Mental Disorder Classification and Diagnosis Standard, third edition; CLZ: = clozapine; DSM-IV: Diagnostic and Statistical Manual of Mental Disorders, fourth edition; DSM-IV-TR: Diagnostic and Statistical Manual of Mental Disorders, fourth edition, text revision; DB: double blind; FD: fixed dosage; HAL: haloperidol; ICD-10: the tenth revision of the International Statistical Classification of Diseases and Related Health Problems; ITT: intent-to-treat; M: monotherapy; NR: not reported; OC: observed cases; OLA: olanzapine; PANSS: Positive and Negative Syndrome Scale; REB: reboxetine; RIS: risperidone; Sz: schizophrenia; SzD: schizophreniform disorder; T: total.

► **Fig. 2** Reboxetine for schizophrenia. Forest plot for clinical efficacy assessed with the PANSS, BPRS, SAPS, or SANS.

found in the Sun et al.'s study (Jadad score < 3 , $p < 0.00001$) but not in the remaining 7 studies (Jadad score ≥ 3 , $p = 0.26$) after leaving out 1 outlying study [27].

Reboxetine administered for 12 weeks reduced weight by a mean of 3.83 kg by reducing appetite and increasing energy expenditure [53]. In a recent RCT of 40 schizophrenia patients, a reboxetine-betahistine combination caused significant weight loss compared to placebo (4.77 vs. 2.02 kg) [54]. In a systematic review of 40 trials, metformin achieved the greatest weight loss (3.17 kg; 95 % CI: -4.44, -1.90) compared to topiramate, sibutramine, aripiprazole, and reboxetine [17]. In another meta-analysis [34] topiramate, aripiprazole, and sibutramine were more effective than reboxetine to induce weight loss. To date, there has been no head-

to-head study or meta-analysis published that directly compared reboxetine and metformin on weight loss.

Adjunctive reboxetine appeared to be safe and well-tolerated. Reboxetine caused more frequent dry mouth (NNH = 14) but less weight gain (NNH = 7) and weight gain of $\geq 7\%$ of the initial weight (NNH = 3). Other ADRs and discontinuation were similar between reboxetine and placebo.

Several studies examined the association between AP-induced weight gain and treatment response in schizophrenia. Treatment response was positively associated with weight gain induced by olanzapine or clozapine [19] and olanzapine and haloperidol [55]. However, only 33 % (3/9) of the RCTs in the present meta-analysis explored the effect of reboxetine on AP-induced weight gain, without

► **Table 2** Subgroup analysis of the associations between moderating variables and change in negative symptoms.

Variables	n (subjects)	SMDs (95 % CI)	I ² (%)	p-value for each subgroup	p-value across subgroups
1. Chinese	3 (341)	-1.06 (-1.29, -0.82)	7	<0.00001	0.0002
Non-Chinese	6 (261)	-0.12 (-0.55, 0.31)	66	0.59	
2. Antipsychotic class: Clozapine	2 (221)	-1.08 (-1.49, -0.66)	53	<0.00001	0.01
Other than Clozapine	7 (381)	-0.27 (-0.74, 0.21)	79	0.27	
3. Trial duration (weeks) ^a : ≥ 12	5 (458)	-0.91 (-1.22, -0.61)	58	<0.00001	<0.00001
< 12	4 (144)	0.16 (-0.19, 0.51)	10	0.37	
4. Age (years) ^a : ≥ 38.0	5 (373)	-0.64 (-1.16, -0.11)	83	0.02	0.37
< 38.0	4 (229)	-0.23 (-0.95, 0.49)	84	0.54	
5. Male predominance (≥ 60 %)	8 (482)	-0.38 (-0.84, 0.07)	82	0.1	0.03
No sex predominance	1 (120)	-1.04 (-1.42, -0.66)	N/A	<0.00001	
6. Study quality: Jadad score ≥ 3	8 (482)	-0.38 (-0.84, 0.07)	82	0.1	0.03
Jadad score < 3	1 (120)	-1.04 (-1.42, -0.66)	0	<0.00001	

^a Analyzed using a mean split method; CI: confidence interval; N/A: not applicable; SMDs: standard mean differences; Bold values: p < 0.05.

► **Fig. 3** Reboxetine for schizophrenia. Forest plot for the change of body weight and BMI.

investigating the association between reboxetine and negative symptoms; thus, the association between AP-induced weight gain and negative symptoms could not be assessed in this meta-analysis.

Three previous meta-analyses [12, 33, 34] of 2 RCTs [18, 22] found similar advantage of adjunctive reboxetine for attenuating weight gain in schizophrenia, but they did not analyze its effects on psychotic symptoms. The current meta-analysis included additional 7 RCTs, allowing more robust and sophisticated analyses, including risk of bias, Jadad scale, GRADE approach, and sensitivity analysis.

There are several limitations of this meta-analysis. First, although all included RCTs were rated as high-quality trials according to the Cochrane risk of bias and the Jadad scale, all 9 RCTs providing the data of co-primary outcomes had relatively small sample size (26–135), which precluded the assessment of publication bias. Second, meta-analyzable results for body weight (I² = 92 %) were heterogeneous, although the statistical significance remained (I² = 0 %) after removing an outlying study from the analysis. Third, reboxetine doses varied across the 9 studies (4.0–8.0 mg/day); therefore, the

dose-response effects of reboxetine in reducing AP-induced weight gain could not be examined. Fourth, cognitive functions were assessed only in 2 trials [26, 27] with conflicting conclusions. The cognitive aspects of reboxetine treatment warrant further investigations. Fifth, all studies had relatively short treatment duration (4–24 weeks); thus, reboxetine's long-term effects on body weight need further studies. Finally, metabolic indices associated with body weight, such as lipid profile, insulin resistance, and leptin could not be analyzed as they were not recorded in the studies.

In conclusion, in this meta-analysis, reboxetine appeared to attenuate AP-induced weight gain in patients with schizophrenia, but it was not consistently effective in treating negative symptoms of schizophrenia. The therapeutic effects of reboxetine on negative symptoms remained doubtful in this meta-analysis as it was driven by an outlying study. In the outlying study [27], schizophrenia patients with metabolic syndrome were recruited. Therefore, the effect of reboxetine on negative symptoms needs to be probably replicated. In addition, high-quality RCTs are warranted to demonstrate reboxetine's long-term safety and efficacy, particularly on cognitive functions.

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Contributors

WZ and XHY selected studies and conducted statistical analysis. XHY and DBC extracted the data. YTX reviewed all the data and helped mediate disagreements. WZ, XBL, and ZMS wrote the first draft. All the authors contributed to the interpretation of data and completing and approving the final manuscript.

Conflict of Interest

The authors declare that they have no conflict of interest concerning this paper.

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