# Mammographic Density and Prediction of Nodal Status in Breast Cancer Patients

## Abstract

**Aim:** Nodal status remains one of the most important prognostic factors in breast cancer. The cellular and molecular reasons for the spread of tumor cells to the lymph nodes are not well understood and there are only few predictors in addition to tumor size and multifocality that give an insight into additional mechanisms of lymphatic spread. Aim of our study was therefore to investigate whether breast characteristics such as mammographic density (MD) add to the predictive value of the presence of lymph node metastases in patients with primary breast cancer.

**Methods:** In this retrospective study we analyzed primary, metastasis-free breast cancer patients from one breast center for whom data on MD and staging information were available. A total of 1831 patients were included into this study. MD was assessed as percentage MD (PMD) using a semiautomated method and two readers for every patient. Multiple logistic regression analyses with nodal status as outcome were used to investigate the predictive value of PMD in addition to age, tumor size, Ki-67, estrogen receptor (ER), progesterone receptor (PR), grading, histology, and multi-focality.

**Results:** Multifocality, tumor size, Ki-67 and grading were relevant predictors for nodal status. Adding PMD to a prediction model which included these factors did not significantly improve the prediction of nodal status (p = 0.24, likelihood ratio test).

**Conclusion:** Nodal status could be predicted quite well with the factors multifocality, tumor size, Ki-67 and grading. PMD does not seem to play a role in the lymphatic spread of tumor cells. It could be concluded that the amount of extracellular matrix and stromal cell content of the breast which is reflected by MD does not influence the probability of malignant breast cells spreading from the primary tumor to the lymph nodes.

## Zusammenfassung


**Methoden:** In dieser retrospektiven Studie sind Patientinnen mit pimären, metastasenfreien Mammakarzinom eingeschlossen wurden, die in einer Institution behandelt wurden. Von allen Patientinnen musste die MD bekannt sein, genauso wie die Staging-Informationen. Die MD wurde semiautomatisiert als prozentuale MD (PMD) von 2 unabhängigen Gutachtern beurteilt. Logistische Regressionsanalysen mit dem Nodalstatus als Zielvariable wurden durchgeführt, um den zusätzlichen prädiktiven Wert von PMD zusätzlich zu Alter, Tumorgröße, Ki-67, Östrogenrezeptorstatus (ER) und Progesteronrezeptorstatus (PR), Grading, Tumortyp und Multifokalität zu bestimmen.

**Ergebnisse:** Multifokalität, Tumorgröße, Ki-67 und Grading waren relevante Prädiktoren des Nodalstatus. Die Hinzunahme von PMD zu diesen Faktoren konnte die Vorhersage des Nodalstatus nicht signifikant verbessern (p = 0.24, Likelihood-Ratio-Test).

**Schlussfolgerung:** Der Nodalstatus konnte relativ gut durch die Parameter Multifokalität, Tumorgröße, Ki-67 und Grading vorhergesagt werden. PMD scheint keine Rolle zu spielen bei der Lymphknotenmetastasierung aus dem Primärtumor. Somit scheint die Menge an Extrazellularmatrix und Stroma, welche das Korrelat der MD...
Introduction

Mammographic breast density (MD) is one of the most important risk factors for breast cancer (BC) [1]. Women with a high MD have a 5-fold increased risk for breast cancer [2–4]. It has been shown that women who develop a high MD after postmenopausal hormone therapy have an increased risk for breast cancer and healthy women who develop a low MD after the intake of tamoxifen have a decreased breast cancer risk [5,6]. Thus, MD is not only a risk indicator but can even serve as a therapy response marker with regard to breast cancer risk. There is limited data on MD and prognosis after the diagnosis of an invasive breast cancer. A large cohort study based in the United States did not find a clear association between mammographic density and prognosis [7]. Only in some subgroups such as women with a high body mass index and women with large tumors, did low MD seem to be associated with a worse prognosis [7]. An analysis of the baseline mammogram in the Swedish mammography screening program yielded a higher risk for breast cancer death in women with a higher MD [8] and a smaller study found a high MD to be an independent predictor of local-regional recurrence after breast-conserving surgery and radiotherapy for invasive breast cancer [9]. Taken together, these results are rather inconsistent and warrant further examination of the correlation of MD with prognosis or other prognostic factors. Additionally, MD has been associated with specific tumor characteristics that are linked to prognosis as well. One study showed an association between progesterone receptor (PR) positivity and MD and between estrogen receptor (ER) beta and MD [10]. Furthermore, it has been reported that MD can increase the risk for ER+/PR+ tumors but not for ER−/PR− tumors [11]. In a case control study MD was associated with the risk for both ER+/PR+ HER2+ and triple negative breast cancer [12]. The largest case control study so far showed that MD increased the risk for high grade tumors and ER negative tumors [13]. We previously could show that a high MD was correlated with ER negativity and PR positivity [14]. Other studies showed different results [15]. Some association has also been reported on the proliferation marker Ki-67. It seems that in some subgroups (low BMI, HRT users and low PR expression) MD was correlated with proliferation in the tumors [16]. Recently it has been shown that a high MD has a histological correlate of high extracellular matrix similar to breast cancer tumors [17]. It could be therefore hypothesized that this might have an influence on tumor spread and lymphogenic metastases in the breast. In fact, a small study showed an association between low MD and lymph node status [18]. Therefore the aim of this study was to examine this correlation in a large case-only study with high quality MD measures, specifically whether MD can improve the prediction of lymph node status in addition to other, previously reported biomarkers that are correlated to a positive lymph node status.

Patients and Methods

Patient selection
The patients were selected from the breast cancer database at the University Breast Center for Franconia. A total of 5,110 patients with invasive BC are documented in the database for the period 1995–2008. In the analysis presented here, patients were excluded in the following hierarchical order: no mammogram performed at the University Breast Center at the time of primary diagnosis or bilateral synchronous breast cancer (excluding 1,989 patients, 485 of whom had bilateral cancer); primary distant metastases at primary diagnosis (excluding 502 patients) unknown nodal status (excluding 788 patients). The final study population consisted of 1,831 patients with incident, unilateral, invasive BC and available MD. The study was approved by the Ethics Committee of the Medical Faculty of Friedrich-Alexander University Erlangen-Nuremberg.

Clinical Data
All patient characteristics and tumor characteristics were documented as part of the certification processes required by the German Cancer Society (Deutsche Krebsgesellschaft) and by the German Society for Breast Diseases (Deutsche Gesellschaft für Senologie) [19] as described before [20–23]. Certification requires tumor characteristics, treatment data, some epidemiological data, histopathological characteristics, tumor treatments, and follow-up to be documented and audited annually.

Histopathology
In the course of routine patient care, dedicated breast pathologists at the University Breast Center for Franconia examined pathologic specimens from all of the patients included in this study. The histological type, grade, resection status, and TNM stage were determined, and expression of ER, PR, and HER2 was analyzed by immunohistochemistry in accordance with standard practice. ER and PR status are based on estimates of the percentage of clearly positively stained nuclei of cancer cells (from 0 to 100%). Ki-67 staining was performed with the MIB1 clone, yielding the percentage of positive nuclei of tumor cells (from 0 to 100%). HER2 testing was performed with the DAKO test (Dako, Denmark). Information about tumor characteristics was subsequently transferred from the histopathological reports our clinical database.

Mammographic Density Measures
The quantitative computer-based threshold density assessments and breast area measurements were made by two different readers with extensive training in the method used. Each mammogram was read by both readers independently of each other. The assessment method has been described and validated previously elsewhere [24]. Briefly, the images (analog images and printouts of processed digital images) were digitized using the CAD PRO Advantage film digitizer (VIDAR, Herndon, Virginia, USA), and for assessment of the density fraction, the reader used the Made- na software program, Version X (Eye Physics, LLC, Los Alamitos, California, USA) [24]. All mammograms were read in random order by two different observers who were unaware of any previous classifications or pathological findings. The average of the
two observers' values for percentage mammographic density (PMD) was used for this analysis, into which only measurements of contralateral cranio-caudal (CC) images were included [36, 37].

Statistical Considerations

Patients were grouped according to their nodal status (negative versus positive). Their characteristics were compared using appropriate unpaired statistical tests. Student's t-tests were used for normally distributed characteristics, Wilcoxon rank-sum tests for ordinal-categorical or continuous non-normally distributed characteristics, and χ² tests for categorical characteristics.

Multiple logistic regression analyses with nodal status as outcome were used to investigate the predictive value of PMD in addition to the well-known predictive factors of age at diagnosis (continuous), tumor size (continuous), Ki-67 (continuous), ER (positive vs. negative), PR (positive vs. negative), grading (ordinal), histology (lobular, ductal, others), and multi-focality (yes vs. no).

Initially, for each continuous predictor, several natural cubic spline logistic regression models, which differ from each other by their degrees of freedom (from 1 to 4), were fitted to explore how the odds ratios (ORs) changed with increasing predictor value [25]. For instance, a cubic spline model with one degree of freedom corresponds to a linear usage of the predictor. The Akaike information criterion (AIC) was applied to choose the best model in order to obtain the suitable degree of freedom for the specific continuous predictor in the following analyses. The AIC was used because it measures goodness of fit and also takes over-fitting into account by penalizing complex models.

The main analysis started with a bootstrap-based selection procedure to identify a set of standard characteristics that were associated with nodal status. Five hundred bootstrap samples of the same size as the data set were selected with replacements. For each bootstrap sample, a backward stepwise variable selection procedure, starting with a logistic regression model with the well-known predictors mentioned above but without PMD, was carried out to obtain the best model in accordance with the AIC. The retained predictors from each bootstrap sample were recorded, and a final variable selection was made by applying a procedure proposed by Sauerbrei and Schumacher [26]. In this procedure the most frequently selected (> 70%) predictors were chosen, and, to address correlation among predictors, the predictor with the larger frequency out of each highly frequent predictor pair (> 90%) was chosen. A logistic regression model with these finally selected predictors was fitted to the original data set (the final model without PMD). Repetitive variable selections were carried out to get a stable stepwise regression result [27].

Next, the bootstrap-based variable selection procedure from above was carried out again, with PMD, the predictors of the final model without PMD, and the interactions between these predictors and PMD, but with the condition that the predictors of the final model without PMD were kept during the stepwise variable selections. Again, a logistic regression model with the finally selected predictors was fitted to the original data set (the final model with PMD).

The final model with PMD was compared with the final model without PMD using the likelihood ratio test. A significant test result means that PMD has a predictive value independent of the well-known predictors considered. The interaction terms of the final model with PMD demonstrate differences within subgroups with regard to the predictive effect of PMD. ORs with 95% confidence intervals based on the final model with PMD were shown.

The predictive ability of the final model with PMD in terms of discrimination and calibration was measured with the area under the receiver operator curve (AUC) and the Hosmer–Lemeshow statistic. The AUC ranges from 0.5 (no discrimination of nodal-positive and nodal-negative patients) to 1 (perfect discrimination). Following Hosmer and Lemeshow, patients were ordered with respect to the predicted probability of positive lymph node status and grouped into equal-sized categories based on percentiles. Frequencies of predicted events were compared with frequencies of observed events using a scatter plot and with the Hosmer–Lemeshow χ² test. A large p-value indicates a satisfactory calibration.

Model building was evaluated by 10-fold cross-validation with 20 repetitions. For this purpose the model-building process was done for each training set, resulting in final models with PMD. These models were fitted to the training sets, and AUCs were calculated with patients of the validation sets. The average AUC was taken as evaluation measure.

All of the tests were two-sided, and a p-value < 0.05 was regarded as statistically significant. Calculations were carried out using the R system for statistical computing (version 2.13.1; R Development Core Team, Vienna, Austria, 2011).

Results

Patient Characteristics

A total of 1831 patients were included in this analysis. Their mean age was 58.9 (± 12.6) years, and 24.0% (n = 429) were pre-menopausal or perimenopausal. The patients' percentage mammographic density was on average 36 (± 19).

Univariate Analysis

Comparing patients with and without involved lymph nodes showed some differences between these groups, most of which were expected. Patients with positive lymph nodes were more likely to have a high proliferative tumor as assessed by Ki-67, a larger tumor size, were more likely to have a multifocal tumor, a ductal or lobular cancer, an ER and PR negative tumor and a higher grading. No differences were seen with regard to PMD and age (Table 1).

Association of PMD and Nodal Status

Preliminary logistic regression analysis showed that the continuous predictors of age at diagnosis, Ki-67, tumor size, and PMD fitted best as cubic spline variables, with 2, 2, 3, and 1 degrees of freedom, respectively. Complete patient characteristics are shown in Table 1.

Multiple logistic regression analysis which did not take PMD into account (see “Patients and Methods”) identified multilocality, tumor size, histology, Ki-67 and grading as relevant predictive factors for nodal status. The predictors age, PR, and ER did not fulfill the final selection criteria, i.e., their predictive value appeared to be irrelevant, or they were already explained by the other factors. Table 2 shows how often each predictor was selected in the bootstrap-based selection process.

The second bootstrap selection process, which took account of the selected predictors and, in addition, PMD and interactions between PMD and those predictors, added PMD but no interaction term (Table 2). The improvement of the prediction model
by PMD was not significant \((p = 0.24, \text{likelihood ratio test})\). The OR for PMD per 10 units increase was 0.96 (95% CI, 0.91 to 1.02). The AUC of the final model with PMD was 0.76, the cross-validated AUC was 0.75. The final model with PMD seemed to be well-calibrated \((p = 0.32, \text{Hosmer-Lemeshow } \chi^2 \text{ test})\). Fig. 1 shows that observed frequencies of lymph node involvement and predicted frequencies by the regression model coincided quite well.

### Discussion

We showed in this retrospective study that mammographic density as measured by computer-assisted assessment of percentage mammographic density did not improve the prediction of nodal status. PMD does not seem to be associated with lymphatic spread of breast cancer cells. Other established variables such as multifocality, tumor size, histology, and proliferation already predicted lymph node status with a high degree of accuracy. Mammographic density correlates to the proportion of radiolucent and radiodense material. While fat tissue is mainly responsible for radiolucency, epithelial cells, fibroblasts, extracellular matrix and other stroma components are responsible for the ab-

### Table 1

<table>
<thead>
<tr>
<th>Characteristic</th>
<th>Value</th>
<th>Node negative (mean or median or N)</th>
<th>SD or interquartile range or %</th>
<th>Node positive (mean or median or N)</th>
<th>SD or interquartile range or %</th>
<th>p-value</th>
</tr>
</thead>
<tbody>
<tr>
<td>Age at diagnosis</td>
<td>years</td>
<td>59.1</td>
<td>12.3</td>
<td>58.5</td>
<td>13.0</td>
<td>0.37</td>
</tr>
<tr>
<td>Ki-67</td>
<td>% positive</td>
<td>15</td>
<td>(5, 25)</td>
<td>20</td>
<td>(10, 40)</td>
<td>&lt;0.000001</td>
</tr>
<tr>
<td>Tumor size</td>
<td>cm</td>
<td>1.4</td>
<td>(1.0, 2.0)</td>
<td>2.4</td>
<td>(1.5, 3.5)</td>
<td>&lt;0.000001</td>
</tr>
<tr>
<td>PMD</td>
<td>%</td>
<td>35</td>
<td>(20, 50)</td>
<td>34</td>
<td>(21, 47)</td>
<td>0.46</td>
</tr>
<tr>
<td>Multifocality</td>
<td>no</td>
<td>906</td>
<td>80.8</td>
<td>437</td>
<td>70.6</td>
<td>&lt;0.000001</td>
</tr>
<tr>
<td>Tumor type</td>
<td>ductal</td>
<td>773</td>
<td>65</td>
<td>443</td>
<td>69.2</td>
<td>&lt;0.001</td>
</tr>
<tr>
<td>lobular</td>
<td>225</td>
<td>18.9</td>
<td>138</td>
<td>21.6</td>
<td></td>
<td></td>
</tr>
<tr>
<td>others</td>
<td></td>
<td>191</td>
<td>16.1</td>
<td>59</td>
<td>9.2</td>
<td></td>
</tr>
<tr>
<td>ER</td>
<td>negative</td>
<td>266</td>
<td>22.3</td>
<td>176</td>
<td>27.5</td>
<td>0.01</td>
</tr>
<tr>
<td>positive</td>
<td>925</td>
<td>77.7</td>
<td>464</td>
<td>72.5</td>
<td></td>
<td></td>
</tr>
<tr>
<td>PR</td>
<td>negative</td>
<td>390</td>
<td>32.7</td>
<td>257</td>
<td>40.2</td>
<td>&lt;0.01</td>
</tr>
<tr>
<td>positive</td>
<td>801</td>
<td>67.3</td>
<td>383</td>
<td>59.8</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Grading</td>
<td>1</td>
<td>189</td>
<td>16.1</td>
<td>35</td>
<td>5.6</td>
<td>&lt;0.000001</td>
</tr>
<tr>
<td>2</td>
<td>769</td>
<td>65.6</td>
<td>417</td>
<td>66.8</td>
<td></td>
<td></td>
</tr>
<tr>
<td>3</td>
<td>214</td>
<td>18.3</td>
<td>172</td>
<td>27.6</td>
<td></td>
<td></td>
</tr>
</tbody>
</table>

### Table 2

Multiple logistic regression analysis, showing frequencies of selected predictors within the first (excluding percentage mammographic density) and second (including percentage mammographic density) bootstrap-based selection process, and inclusion (yes/no) in the final logistic regression models according to pre-specified criteria. (PMD: percentage mammographic density; ER: estrogen receptor; PR: progesterone receptor).

<table>
<thead>
<tr>
<th>Predictor</th>
<th>1st selection procedure</th>
<th>2nd selection procedure</th>
</tr>
</thead>
<tbody>
<tr>
<td>Multifocality</td>
<td>100</td>
<td>yes</td>
</tr>
<tr>
<td>Tumor size</td>
<td>100</td>
<td>yes</td>
</tr>
<tr>
<td>Histology</td>
<td>92</td>
<td>yes</td>
</tr>
<tr>
<td>Ki-67</td>
<td>81</td>
<td>yes</td>
</tr>
<tr>
<td>Grading</td>
<td>77</td>
<td>yes</td>
</tr>
<tr>
<td>Age at diagnosis</td>
<td>58</td>
<td>no</td>
</tr>
<tr>
<td>PR</td>
<td>35</td>
<td>no</td>
</tr>
<tr>
<td>ER</td>
<td>25</td>
<td>no</td>
</tr>
<tr>
<td>PMD</td>
<td>-</td>
<td>-</td>
</tr>
<tr>
<td>PMD × Ki-67</td>
<td>-</td>
<td>-</td>
</tr>
<tr>
<td>PMD × tumor size</td>
<td>-</td>
<td>-</td>
</tr>
<tr>
<td>PMD × multifocality</td>
<td>-</td>
<td>-</td>
</tr>
<tr>
<td>PMD × grading</td>
<td>-</td>
<td>-</td>
</tr>
<tr>
<td>PMD × histology</td>
<td>-</td>
<td>-</td>
</tr>
</tbody>
</table>

<sup>a</sup> PMD and any interaction terms (e.g., PMD x Ki-67) were not considered in the 1st selection process.

<sup>b</sup> The finally selected predictors in the first selection process (“yes” in column 3) were kept in all bootstrap samples of the second selection process.

<sup>c</sup> Predictors not included in the final model without PMD (“no” in column 3) were not regarded in the second selection process.
In conclusion, we could not show mammographic density to be associated with the spread of breast cancer tumor cells to the axillary lymph nodes. As mammographic density is mainly a correlate of the amount of extracellular matrix and stroma, the presence or absence of these components and the associated molecular biology do not seem to play a major role in lymphatic spread. Other factors such as tumor size and multifocality already explained the variability of lymph node metastasis to a high degree in our study.

Acknowledgement
Katharina Heusinger was in part funded by the ELAN program (“Erlanger Leistungsbezogene Anschubfinanzierung und Nachwuchsförderung”) of Erlangen University Hospital.

Conflict of Interest
The authors declare that they do not have a conflict of interest.

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