Factors for Preterm Births in Germany – An Analysis of Representative German Data (KiGGS)

Einflussfaktoren auf die Frühgeburt in Deutschland – Analyse der für Deutschland repräsentativen KiGGS-Daten

Abstract

Introduction: Preterm birth is a global scourge, the leading cause of perinatal mortality and morbidity. This study set out to identify the principal risk factors for preterm birth, based on the German Health Interview and Examination Survey for Children and Adolescents (KiGGS). A range of possible factors influencing preterm birth were selected for inclusion in the questionnaire, covering factors such as gender, national origin, immigrant background, demography, living standard, family structure, parental education and vocational training.

Methods: All data were taken from the aforementioned KiGGS survey conducted between 2003 and 2006. A total of 17641 children and adolescents (8656 girls and 8985 boys) drawn from 167 German towns and municipalities deemed to be representative of the Federal Republic of Germany were included in the study. Gestational age at birth was available for 14234 datasets. The questionnaire included questions from the following areas as possible factors influencing preterm birth: gender, national origin, immigrant background, demography, living standard, family structure, parental education and vocational training.

Results: The preterm birth rate was 11.6%, higher than that of other national statistical evaluations. Around 57.4% of multiple pregnancies and 10% of singleton pregnancies resulted in preterm delivery. Multiple pregnancy was found to be the most important risk factor (OR 13.116). With regard to national origins and immigration background, mothers from Turkey, the Middle East, and North Africa had a higher incidence of preterm birth. Preterm birth was more prevalent in cities and large towns than in small towns and villages.

Conclusion: Risk factors associated with preterm birth were identified. These should help with the early identification of pregnant women at risk.

Zusammenfassung

Einleitung: Die Frühgeburt ist die Hauptursache für perinatale Mortalität und Morbidität und ein weltweites Problem. Das Ziel dieser Studie war, die Hauptrisikofaktoren für eine Frühgeburt aus dem Kinder- und Jugendgesundheits-Survey (KiGGS) zu identifizieren. Als mögliche Einflussfaktoren auf die Frühgeburt wurden im Fragenkatalog Fragen aus den Bereichen Geschlecht, Herkunft, Migrationshintergrund, Demografie, Lebensumstände, Familienstrukturen, Schul- und Berufsausbildung der Eltern ausgewählt.


Zusammenfassung: Risikofaktoren, die mit einer Frühgeburt assoziiert waren, wurden identifiziert. Diese sollen helfen, Risikoschwangere rechtzeitig zu identifizieren. Die Frühgeburtsrate
The preterm birth rate in our survey was higher than that found in other national statistical evaluations based on process data. More than half of all multiple pregnancies ended in preterm birth.

Introduction

Preterm birth (defined as birth before 37 + 0 weeks of pregnancy) is the principal cause of perinatal morbidity and mortality. Over the last two decades the preterm birth rate has remained unchanged or even risen in most countries, despite the increased understanding of possible risk factors and their pathological mechanisms [1–5]. In 2013 the preterm birth rate for Germany stood at 8.7% [6], one of the highest recorded in Europe [7]. Globally, the rate varies between 7–14%, which amounts to some 15 million babies born prematurely every year. Approximately 75% of all cases of perinatal mortality and more than half of all cases of long-term perinatal morbidity are associated with premature delivery [8]. With an incidence of 27%, preterm delivery is the principal cause worldwide of neonatal death (occurring within the first 28 days of life), which translates into more than one million dead children every year [9]. Typical secondary complications associated with preterm birth include neonatal respiratory distress syndrome (RDS), persistent ductus arteriosus (PDA), retinopathy of prematurity (ROP), intraventricular hemorrhage (IVH), periventricular leukomalacia (PVL), and bronchopulmonary dysplasia (BPD) [10].

Preterm birth-related mortality and morbidity decreased following the introduction of antenatal corticosteroid therapy and improved neonatal intensive care. However, in addition to the abovementioned immediate perinatal outcomes, there are other consequences that severely affect the subsequent lives of affected children and their families. Typical sequelae affecting preterm born children in later life are cerebral palsy and respiratory distress syndrome, along with sensory deficits and learning difficulties [11]. In addition, the height of children born prematurely is often smaller than average in later life [12].

The reproductive phase of women who were themselves born prematurely is also affected. Thus, a low maternal birth weight (irrespective of whether this was caused by preterm birth or growth restriction) is not only a risk factor for hypertensive (including pre-eclampsia) and diabetogenic pregnancy complications [13, 14], but also for preterm birth in the next generation [15]. Improved survival rates have also affected the decisions taken by physicians when weighing up the decision to continue a pregnancy against the risks associated with preterm delivery, the consequence being an increased number of iatrogenic preterm births [16].

The causes of preterm birth can be classified as follows:
1. Premature delivery for fetal or maternal indications (e.g. pre-eclampsia, placenta previa bleeding, multiple pregnancies, or fetal growth restriction),
2. Spontaneous preterm labor without rupture of membranes, and
3. Premature rupture of membranes resulting in preterm delivery.

Some 70% of cases can be assigned to the latter two groups [17, 18], referred to here as “spontaneous preterm birth.” Complex mechanisms seem to underpin the etiology; data from previously published studies suggest that infections such as cho-

Risk factors also include emotional stress and other psychological stressors. Thus it was found that a number of pregnant women suffering from posttraumatic stress disorder in the wake of the 9/11 attacks of 2001 went on to give birth prematurely (OR 2.48, 95% CI 1.05–5.84) [25].

Causal connections between the majority of these risk factors and preterm birth are, however, difficult to prove, since it has not been possible to identify these connections in randomized and controlled studies. Since the German Health Interview and Examination Survey for Children and Adolescents (KiGGS) is a cross-sectional study, it only describes associations and prevalences. However, causes and consequences can often be plausibly assigned based on prior knowledge. A high degree of validity can be assumed if:
1. data were acquired in double-blind studies,
2. the significance level is less than p < 0.01,
3. dose-response relationships are present, and
4. results do not vary, even when taking cofactors into consideration.

Cross-sectional studies cannot provide proof of etiology; other study designs are required for that [26, 27]. The present study aimed to investigate the following issues:
1. Is it possible to determine preterm birth rates from the data available from the German Health Interview and Examination Survey for Children and Adolescents (KiGGS)?
2. Which children are affected by preterm birth? Is it possible to identify potential risk factors for preterm birth in Germany? The aim of our study was therefore to identify high-risk pregnancies based on a knowledge of the risk factors with a view to improving prenatal care. Preventive measures can be taken early on if there are clinical indications that preterm birth is imminent.

Methods

Selection of participants and database

Our data were obtained from the German Health Interview and Examination Survey for Children and Adolescents (KiGGS) study on child and adolescent health in Germany. Commissioned by the Robert Koch Institute, the KiGGS study was conducted from 2003 to 2006, making it the first of its kind to be attempted in Germany. Certainly it was the first to compile data that were representative for the country as a whole and could be used to make
general assumptions about the health of children and adolescents between the ages of 0–17 [28–30].

Briefly, participants were chosen using a two-step procedure: the first step consisted of selecting 167 sample points deemed representative for Germany, in that they reflected the population-specific extent of urbanization and geographic distribution. To ensure that sample sizes were adequate and the data could be generalized according to place of residency (new federal states, old federal states, Berlin), the breakdown for the 167 data collection sites was as follows: 112 were in the old federal states, 50 in the new federal states, with the remaining 5 in Berlin. Then, in a second step, persons were selected from the population registers kept by the local authorities, using an age- and gender-based random sampling procedure.

For each and every data collection site, a number of children (8, 9 or 10, depending on the size of the community) were chosen for each birth cohort, resulting in (respectively) 144, 162 and 180 persons deemed suitable for study who were then asked to participate.

Auxiliary data on targeted risk factors and covariables were obtained from questionnaires sent to parents and their children. The participation rate was 66.6%, and varied only slightly between age groups and genders. Participation rates differed between immigrants and Germans, between cities and large towns (more than 100 000 inhabitants) and other data collection sites with fewer inhabitants, but also between the old and the new federal states (including Berlin). The percentage of quality-neutral data losses was comparatively low (5.3%).

With a view to securing representative findings, analyses were adjusted using a weighting factor designed to correct for deviations from the demographic structure (as on December 31, 2004) in terms of age, gender, region (Berlin, East and West Germany) and nationality.

For the background methodologies informing the KIGGS study and other particulars, readers should consult other publications [31–33].

**Study population**

A total of 17641 children and adolescents (8656 girls and 8985 boys) were studied from 167 cities, towns, and villages, selected as representative for the Federal Republic of Germany. The study population comprised 14836 children and adolescents over the age of three. There was an oversampling of participants with an immigrant background and also of participants from the new federal states.

The data for analysis were obtained from parents in the form of written statements (questionnaires); maternal data was obtained from prenatal care records (the “maternal health passport”, issued to all expectant mothers in Germany documenting prenatal and natal care) and additional data was obtained from pediatric healthcare records (yellow pediatric medical checkup booklet). Subsequent steps included physical examinations, anthropometry, and laboratory tests. Blood samples were collected from 86% of the persons studied, after first securing a positive vote from the appropriate ethics commission since it was expected that each child would benefit [34].

**Statistical analysis**

Statistical analysis (probability values and confidence intervals) was performed using complex sampling techniques in SPSS, Version 20 (IBM, Armonk, NY, USA, www.ibm.com). Differences between groups with probability values (p-values) of less than 0.05 or 0.01 or with non-overlapping confidence intervals of 95% were considered statistically significant.

To start with, datasets where the gestation age at delivery was known were identified and the preterm births were selected. Questions focusing on possible factors influencing preterm birth such as gender, nationality, immigrant background, demography, living standards, family structure, parental education and vocational training, parental occupational status, and household income were chosen; the responses were analyzed using bi- or multivariate statistical correlation analysis.

Stratified prevalences were determined and the probabilities for differences between groups were calculated using chi² test for nominal data and Kendall’s tau b correlation coefficient for ordinal data.

The risk factors for preterm birth were determined using multivariable logistic regression, and odds ratios (OR) and 95% confidence intervals (95% CI) were specified. Linear regression equations were used to show changes in birth rates over time.

**Results**

A total of 17641 children and adolescents (8656 girls and 8985 boys) were studied from 167 cities, towns, and villages, chosen as representative for Germany. Gestational age at birth was known for 14234 datasets. **Table 1** shows the key data on risk factors for preterm birth. The datasets were analyzed using bi- or multivariate statistical correlation analysis. Because of the complexity, correlations and associations are listed according to the specific issue.

**Preterm birth rate**

With regard to the preterm birth rate, 11.6% of births occurred before the 38th week of gestation (< 37 + 0 GW).

In the annual birth cohorts (1985–2006), figures ranged between 8 and 14.9%. The gender distribution for preterm births was 12% (for boys) and 11.1% (for girls) (p = 0.047).

57.4% of multiple births and 10% of singleton births were premature. Over time, the preterm birth rate showed a clear upward trend (Fig. 1).
Parental nationality

The preterm birth rate was significantly higher for immigrants than non-immigrants (13.8 vs. 11.2%, p < 0.001). German mothers had a preterm birth rate of 11%, whereas the figures for mothers from Central Europe, Black Africa and Latin America were respectively 12.6%, 13.6%, and 15.3%. Mothers from Turkey, the Near East, the Middle East and North Africa (17.3%) and from Asia (17.2%) (p < 0.001) were most likely to have preterm births.

11.1% of families with a German father had a preterm birth (as opposed to 10.1% for families with a European father and 11.6% for the overall population). This is in contrast to the figures for fathers from Turkey, the Near East, the Middle East and North Africa (16%), sub-Saharan Africa (16%), Latin America (16.2%) and Asia (16.3%) (p < 0.001).

For immigrant mothers, the preterm birth rate depended on the duration of residence in Germany. For mothers who had only arrived in Germany in the year of delivery, the preterm birth rate was 12.4%, whereas the rate for women who had resided in Germany between two and ten years increased to 14.3% before declining again to 11.3% (overall population: 11.6%; p = 0.003). The paternal duration of residence was not statistically significant.

Effect of place of residency on preterm birth rate

The preterm birth rate correlated to a highly significant degree with municipal size. Whereas for rural regions the preterm birth rate was 9.5%, the corresponding figure for small towns was 11.7%; for midsized towns it was 12.2%, and for large towns and cities it was 12.1% (p = 0.001). In comparison, there was no significant difference (p = 0.301) in preterm birth rates between the new (12.1%) and old federal states (11.4%).

Table 1  Logistic regression for risks underlying preterm birth. OR (Odds Ratio), 95% CI (95% confidence interval), p-value (significance). Significant results are indicated in bold.

<table>
<thead>
<tr>
<th>Risk factor*</th>
<th>OR **</th>
<th>95% CI**</th>
<th>p-value</th>
</tr>
</thead>
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<tr>
<td></td>
<td>lower</td>
<td>upper</td>
<td></td>
</tr>
<tr>
<td>Pregnancy factors</td>
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<td></td>
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<td>13.116</td>
<td>10.896</td>
<td>15.790</td>
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<tr>
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<td>1.012</td>
<td>0.526</td>
<td>1.944</td>
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<td>1.117</td>
<td>1.010</td>
<td>1.234</td>
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<td>Nullipara</td>
<td>1.470</td>
<td>1.322</td>
<td>1.634</td>
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<td>20–29</td>
<td>0.983</td>
<td>0.750</td>
<td>1.287</td>
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<td>30–39</td>
<td>1.182</td>
<td>0.891</td>
<td>1.567</td>
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<td>40 and above</td>
<td>1.212</td>
<td>0.745</td>
<td>1.971</td>
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<tr>
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<td>0.775</td>
<td>0.323</td>
<td>1.860</td>
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<td>Maternal height (per 10 cm)</td>
<td>0.804</td>
<td>0.746</td>
<td>0.867</td>
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<td>&lt; 19 kg/m²</td>
<td>1.315</td>
<td>1.057</td>
<td>1.637</td>
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<tr>
<td>24–&lt; 26 kg/m²</td>
<td>0.879</td>
<td>0.759</td>
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<td>26–&lt; 28 kg/m²</td>
<td>0.952</td>
<td>0.798</td>
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<td>28–&lt; 30 kg/m²</td>
<td>0.764</td>
<td>0.613</td>
<td>0.951</td>
</tr>
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<td>30–&lt;35 kg/m²</td>
<td>0.799</td>
<td>0.654</td>
<td>0.977</td>
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<td>≥ 35 kg/m²</td>
<td>1.109</td>
<td>0.865</td>
<td>1.421</td>
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<tr>
<td>middle</td>
<td>0.805</td>
<td>0.713</td>
<td>0.910</td>
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<tr>
<td>high</td>
<td>0.701</td>
<td>0.606</td>
<td>0.812</td>
</tr>
<tr>
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<td>1.168</td>
<td>1.021</td>
<td>1.335</td>
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<tr>
<td>small town</td>
<td>1.239</td>
<td>1.056</td>
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<td>1.315</td>
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<td>1.202</td>
<td>1.019</td>
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<td>Europe/North America/Australia</td>
<td>1.112</td>
<td>0.943</td>
<td>1.310</td>
</tr>
<tr>
<td>Turkey/Near East/Middle East/North Africa</td>
<td>1.501</td>
<td>1.207</td>
<td>1.867</td>
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<tr>
<td>Sub-Saharan Africa</td>
<td>1.292</td>
<td>0.510</td>
<td>3.272</td>
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<tr>
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<td>1.262</td>
<td>0.584</td>
<td>2.726</td>
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<td>0.996</td>
<td>2.197</td>
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<tr>
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<td>0.580</td>
<td>1.746</td>
</tr>
<tr>
<td>Constant</td>
<td>3.966</td>
<td></td>
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</tr>
</tbody>
</table>

* Reference categories are not listed, e.g. singleton birth, girl, multipara (child with older siblings), country, old federal states, rural area, low social status, Germany, under 20 years of age, 19–<24 kg/m².

** An OR of more than 1 indicates a higher risk; it is significant when the upper and lower 95% CI are over 1. An OR of less than 1 indicates a lower risk. It is significant when the 95% CI is always less than 1.
Social status and preterm birth

Social status was determined based on parental statements on
1. education and vocational training,
2. household income, and
3. occupational status.

Following Winkler and Lange [35,36], social status was calcu-
lated as an index based on the above-listed indicators. To this
time, each of the three indicators was assigned a value ranging
from 1–7; these values were then added together to create a
spectrum which ranged from a minimum of 3 points to a maxi-
imum of 21 points. This permitted a family’s social status to be
classified, with 3–8 points indicating a low status, 9–14 points a
middle social status, and 15–21 points a high social status.

Occupational status for the purposes of classification was defined
as the current job held or, failing that, the last job held.

Taking social status into consideration resulted in significant re-
results. The preterm birth rate was found to be 13.5% for low, 11.1%
for middle, and 10.5% for high social status groups (p < 0.001).

Multivariable analysis of risks associated
with preterm birth

When logistic regression was used to analyze the risk for preterm
birth, the respective risk factors were added incrementally (Table
1).

Maternal age referred to age at the time of delivery; however,
body mass index (BMI) only referred to BMI at the time of the
survey, which may not adequately reflect the association at the
time of delivery.

Multiple pregnancy was the most significant risk factor for pre-
term birth. The risk was, in fact, 13 times higher than for single-
ton pregnancies (Table 1).

This was followed by parity; however, the number of prior preg-
nancies or of older siblings was not significant, prompting the de-
cision to take both figures together. Thus, the risk for nulliparae
was found to be 1.5 times higher (Table 1).

Children from the new federal states had an only marginally
higher risk (less than 20%) of preterm birth (Table 1).

A further maternal risk factor studied was height, stratified in in-
crements of 10 cm to simplify presentation. Mothers from towns,
especially those of medium height, had a higher risk of preterm birth
than mothers from rural areas. Middle and high social sta-
tus reduced risk, whereas an immigrant background increased it
(Table 1). Maternal age did not play a significant role. Risk de-
creased with increased maternal height and was highest for low
or very high BMI (Table 1).

Discussion

Preterm birth is a global scourge. A rise in preterm birth rates has
been observed in almost all countries all over the world. Preterm
birth is the leading cause of perinatal mortality. Every year, an es-
timated 15 million babies are born prematurely worldwide, and
more than one million children die annually from the sequelaes
of preterm birth [2]. In industrially more developed countries
the preterm birth rate is 7.5%; the corresponding figures for less
and least developed countries are 8.8% and 12.5%, respectively
[37].

Chorioamnionitis is an important cofactor for preterm birth. Hist-
ologically verified chorioamnionitis is found in approx. 60–80%
of the placasas of mothers who gave birth prior to the 28th
week of gestation, as well as in 40–50% of the placasas of moth-
ers who gave birth between the 29th and 34th weeks of gestation
[19].

The World Health Organization (WHO) classifies preterm births
into three categories; “extremely preterm” (< 28 GW); “very pre-
term” (28 to < 32 GW) and “moderate to late preterm” (32 to < 37
GW).

“Late preterm infants” are the most rapidly growing subgroup of
preterm births and account for more than 80% of all preterm
births (Germany 2013: 83% [6]).

In addition to the direct consequences, preterm birth also carries
an increased risk of late sequelae, such as chronic conditions and
higher mortality rates in later life. Many survivors suffer from the
consequences, which may include learning difficulties as well as
visual and auditory impairment, for the rest of their lives. In ad-
inion to pneumonia, complications associated with preterm
birth are among the primary causes of death in children under
the age of five [38].

According to the figures issued for 2013 by the Institute for Ap-
plied Quality Improvement and Research in Healthcare GmbH
(Institut für angewandte Qualitätsförderung und Forschung im
Gesundheitswesen; AQUA), the preterm birth rate for all of Ger-
many was 8.7% [6].

However, this is lower than the preterm birth rate calculated from
the KiGGS data, which stands at 11.6% for the overall popu-
lation. When the birth years were considered individually (e.g.
1985–2006), the preterm birth rate ranged between 8 and
14.9%. For the period 1994–2006, it was possible to compare the
data from the KiGGS survey with the preterm birth rates from the
German Federal Health Reporting system (Gesundheitsbericht-
erstattung des Bundes, GBE Bund) (Fig. 1). The data of the
KiGGS survey differed significantly from the data supplied by
GBE Bund (p = 0.0002). In fact, for every year of observation, the
KiGGS data was approx. 2–3 times higher than the corresponding
figures from GBE Bund, even though the trend lines were virtu-
ally parallel for both datasets (a rise in the KiGGS trend line of
0.0031 vs. a rise in the GBE trend line of 0.0026).

This is particularly interesting because surveys like KiGGS are
recognized as offering the best method to determine prevalence
due to their direct and representative acquisition of data from in-
dividuals.

A possible cause of this discrepancy could be due to the fact that,
at the time, there was no systematic, structured perinatal docu-
mentation system across all of Germany, meaning that it was dif-
cult to compare data of individual federal states [39].

Another explanation could be the quality of the process data, i.e.
data relating to the care process. Non uniform documentation
methods in delivery wards could have skewed data outcomes.

Comparisons of preterm birth rates from different decades
should be treated with caution as it is by no means certain that
the data are comparable. As survival rates for extremely prematu-
ture and extremely low birthweight neonates improved, registra-
tion procedures may have changed over time.

With regard to the incidence of multiple pregnancies, earlier
studies have shown that in all industrial countries multiple birth
rates are trending upward, chiefly due to the introduction of re-
productive medical techniques [40]. Some 40–60% of all multiple
births are premature [24]. Perinatal mortality is four times higher
for twins than for singleton births; with triplets the perinatal
mortality is six times higher [41]. Morbidity and mortality of
twins born as a result of reproductive medical techniques do
not, however, diverge from those of spontaneously conceived
twins [42].
According to Hellin’s rule, the frequency of spontaneous multiple births is depicted by the formula $1:85^n − 1$, with $n$ corresponding to the number of children born. Calculated for twins, the frequency is $1:85$ (1.2%); for triplets the figure is $1:85^2 = 1:7225$ (0.014%). A higher prevalence of multiple pregnancies was also found in our analysis. In the observation period covered by the KiGGS survey, the percentage of children resulting from multiple pregnancies averaged 3.3%, exhibiting a slightly upward trend. These results are consistent with statistical data from the German Federal Statistical Office [43]. See Fig. 2 for a comparison of the two datasets.

A multiple pregnancy is a significant risk factor for preterm birth. Our data showed that 57.4% of multiples were born prematurely (Fig. 3). Trend analysis showed an almost tenfold rise over time for multiples as opposed to singletons (0.0019 vs. 0.0183). These results are consistent with previously published data. In their analysis of data from Berlin, Bergmann et al. noted an increase from 1.3 to 1.8% in multiple births between 1993 and 1999. This study observed an increase in the proportion of multiple births as well as in the number of pregnancies following reproductive medical treatment [44].

Where maternal BMI is concerned, it is known that extremes of weight, i.e. obesity and very low maternal weight, are correlated with an increased risk of preterm birth. Despite the simultaneous consideration of other factors that affect BMI in the logistical model such as low socioeconomic status or ethnicity, the correlation persists [45, 46]. Both women with a BMI of < 19 kg/m² and those with a BMI > 35 kg/m² had an increased risk of preterm birth (OR 1.315 and 1.109 respectively).

With regard to the role played by parental origins, it could be shown that an immigrant background was a risk factor for preterm birth. Mothers from Turkey, the Near East, the Middle East, North Africa, and Asia (grouped together here due to similar risk patterns) were most commonly affected. One possible explanation might be that immigrants from countries in these parts of the world had to leave their homelands under extremely stressful conditions (political tensions, war). In a Swedish study by Liu et al., the risk of preterm birth among immigrants from countries wracked by civil war was evaluated. It was found that the risk was greatest in the first year of residency in Sweden (1.39), but declined after a period of two or three years [47].

A correlation between the duration of stay in Germany and the preterm birth rate was also evident in our own analysis (year 1: 12.4%; years 2–10: 14.3%; after 11 years: 11.3%). In this context, mention should also be made of the so-called “Latina epidemiologic paradox” also known as the “healthy migrant effect,” which has been observed in almost all immigrant countries. As originally described, the paradox was that Americans of Hispanic or Latin American origin (Latinas) exhibited a rather better health profile than Americans of Caucasian origin, despite Latinas often having a lower socioeconomic status [48]. In terms of pregnancies of Mexican women, the paradox was that they were less likely to give birth to underweight babies or have preterm deliveries than Americans of non Mexican origin; the decreased risk factor was in the order of 30% [49]. This paradox was only found, however, in the first generation of immigrants. Immigrants, for one thing, arrive with certain traditions and habits that promote healthy outcomes. They may have, for example, suppor-
tive family structures to draw on in their new country, structures that play out positively in terms of reproductive health [50, 51]. In the course of adapting to an unfamiliar environment, so called acculturation processes, these health-promoting habits were no longer documented for second and third generation immigrants [52].

The figures for our study show that after ten years of residence the preterm birth rate of immigrants achieved parity with that of non immigrants (11.3 vs. 11.6%). Evidently, this is a highly complex issue which is influenced by such diverse factors as country of origin, circumstances of arrival in the new country (classic immigration vs. refugees fleeing war), social status, and access to medical care.

Urban women were more likely to give birth prematurely than rural women. Our data showed, moreover, that while women of immigrant background living in larger communities were at greater risk of preterm birth, the preterm birth rate declined the longer they lived in Germany.

These results are in accordance with previously published studies. Thus, for instance, Kent et al. showed that inhabitants of urban areas are more likely to miscarry or give birth prematurely. As exemplified by Afro-American women living below the poverty line, it could be shown that urban women had a greater risk of miscarriage or preterm delivery than their rural counterparts [53].

Social status is an important cofactor in the incidence of preterm birth. Several authors reported that women of low social status are more likely to suffer preterm birth [53, 54]. In a study by Straube et al., the preterm birth rate for unskilled and semiskilled working women was 7.8%, whereas the corresponding figure for women working as senior civil servants was lower (6.3%). When combined with other risk factors, the figure for the former even increased exponentially. Thus, the preterm birth rate for working women aged ≥40 years with more than four living births jumped from 8.2 to 13.8%; in contrast, the figure for senior civil servants in the same category was 5.6% [55].

The importance of social status is further emphasized by the fact that higher social status not only lowers the risk of preterm birth, it also reduces the incidence of cognitive and social deficits in later life among children born prematurely. Treyvauad et al. showed, for instance, that an optimal home environment is conducive to prematurely born babies developing better cognitive and socio-emotional skills in infancy [56]. Our own results confirmed the role played by social status. The OR for a preterm birth for mothers of medium and high social status were 0.805 and 0.701, respectively.

**Conclusion**

In summary, the present study is the first systematic and generalizable analysis of the risk factors for preterm birth in Germany. The preterm birth rate in the KiGGS survey was 11.6%, but when stratified according to year of birth (1985–2006), it rose up to 14.9%.

Remarkably, the preterm birth rate in the present survey was significantly higher than that of either GBE Bund or the German Federal Statistical Office, both of which are based on process data. Despite intensive efforts, it has not so far proved possible to prevent or reduce preterm birth. Neither comprehensive information on known risk factors nor the better availability of medical care during pregnancy nor the proximity of well equipped perinatal care centers nor a profusion of therapeutic approaches (to-
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